

Summary of Estimators

Avadhesh Yadav(Y8104010)

Vishesh Lokras(Y8104067)

Guided by:

Dr. R. Hegde

Approaches to Estimation

- Classical Approach
 - Unknown parameter vector θ - ***Deterministic Constant***
 - Cramer - Rao Lower Bound
 - Rao-Blackwell-Lehmann- Scheffe
 - Best Linear Unbiased Estimator
 - Maximum Likelihood Estimator
 - Least Squares Estimator
 - Method of Moments
- Bayesian Approach
 - Known parameter vector θ - ***Realization of a Random Vector***
 - Minimum Mean Square Error Estimator
 - Maximum A Posteriori Estimator
 - Linear Minimum Mean Square Error Estimator

Estimator Overview

- Data Models
- Estimator
- Optimality/Error Criterion
- Performance
- Comments etc..

Cramer Rao Lower Bound(CRLB)

a. Data Model/Assumption

PDF $p(\mathbf{x};\boldsymbol{\theta})$ is known.

b. Estimator

If the equality condition for the CRLB

$$\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta})$$

is satisfied, then the estimator is $\hat{\boldsymbol{\theta}} = \mathbf{g}(\mathbf{x})$

where $\mathbf{I}(\boldsymbol{\theta})$ is a $p \times p$ matrix dependent only on $\boldsymbol{\theta}$ and $\mathbf{g}(\mathbf{x})$ is a p -D function of data \mathbf{x} .

c. Optimality/Criterion

$\hat{\boldsymbol{\theta}}$ **achieves the CRLB**, the lower bound on the variance for any unbiased estimator(and hence is said to be efficient), and is therefore the MVU estimator.

d. Performance

It is **unbiased** and has the **minimum variance**

$$\text{var}(\hat{\theta}_i) = [\mathbf{I}^{-1}(\boldsymbol{\theta})]_{ii} \quad i=1,2,\dots,p$$

where

$$[\mathbf{I}(\boldsymbol{\theta})]_{ij} = E\left[\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \theta_i} \frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \theta_j}\right]$$

e. Comments

An efficient **Estimator may not exist** & this approach may fail

Rao-Blackwell-Lehmann-Scheffe(RBLS)

a. Data Model/Assumption

PDF $p(\mathbf{x};\boldsymbol{\theta})$ is known.

b. Estimator

i. Find a *sufficient statistic* $\mathbf{T}(\mathbf{x})$ by factoring PDF as

$$p(\mathbf{x};\boldsymbol{\theta})=g(\mathbf{T}(\mathbf{x}),\boldsymbol{\theta})h(\mathbf{x})$$

ii. If $E[\mathbf{T}(\mathbf{x})]=\boldsymbol{\theta}$, then $\hat{\boldsymbol{\theta}}=\mathbf{T}(\mathbf{x})$. If not, we must find a p -dimensional function \mathbf{g} so that $E[\mathbf{g}(\mathbf{T})]=\boldsymbol{\theta}$, and then $\hat{\boldsymbol{\theta}}=\mathbf{g}(\mathbf{T})$.

c. Optimality/Criterion

$\hat{\boldsymbol{\theta}}$ is the MVU estimator.

d. Performance

Unbiased but **variance depends upon PDF** -no general formula is available

e. Comments

“**Completeness**” of sufficient statistic must be checked. A p -dimensional sufficient statistic may not exist so this method may fail.

Best Linear Unbiased Estimator

a. Data Model/Assumptions

$$E(\mathbf{x}) = \mathbf{H}\boldsymbol{\theta}$$

where \mathbf{H} is an $N \times p$ ($N > p$) known matrix and \mathbf{C} , the covariance matrix of \mathbf{x} , is known. Equivalently, we have

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

where $E(\mathbf{w}) = 0$ and $\mathbf{C}_w = \mathbf{C}$.

b. Estimator

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x}$$

c. Optimality/Error Criterion

$\hat{\boldsymbol{\theta}}_i$ for $i = 1, 2, \dots, p$ has **minimum variance** of all unbiased estimators that are *linear* in \mathbf{x} .

d. Performance

$\hat{\boldsymbol{\theta}}_i$ for $i = 1, 2, \dots, p$ is unbiased. The variance is

$$\text{var}(\hat{\theta}_i) = [(\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}]_{ii} \quad i = 1, 2, \dots, p$$

e. Comments

If \mathbf{w} is a Gaussian RV, then is $\hat{\boldsymbol{\theta}}$ also the **MVU estimator** (for all linear & non linear functions of \mathbf{x}).

Maximum Likelihood Estimator

a. Data Model/Assumptions

PDF $p(\mathbf{x};\boldsymbol{\theta})$ is known

b. Estimator

$\hat{\boldsymbol{\theta}}$ is the value of $\boldsymbol{\theta}$ maximizing $p(\mathbf{x};\boldsymbol{\theta})$, where \mathbf{x} is replaced by the observed data samples.

c. Optimality/Error Criterion

Not optimal in general. Under certain conditions on the PDF, however, the MLE is efficient for large data records or as $N \rightarrow \infty$ (asymptotically). Hence, asymptotically it is the MVU estimator.

d. Performance

For finite N depends on PDF-no general formula is available. Asymptotically, under certain conditions

$$\hat{\boldsymbol{\theta}} \square N(\boldsymbol{\theta}, I^{-1}(\boldsymbol{\theta}))$$

e. Comments

If an MVU estimator exists, the MLE procedure will produce it .

Least Squares Estimator

a. Data Model/Assumption

$$x[n]=s[n;\boldsymbol{\theta}] + w[n] \quad n=0,1,\dots,N-1$$

where the signal $s[n;\boldsymbol{\theta}]$ depends explicitly on the unknown parameters. Equivalently, the model is

$$\mathbf{x}=\mathbf{s}(\boldsymbol{\theta}) + \mathbf{w}$$

where \mathbf{s} is a known N-dimensional function of $\boldsymbol{\theta}$ and the noise or perturbation \mathbf{w} has zero mean.

b. Estimator

$\hat{\boldsymbol{\theta}}$ is the value that minimizes

$$J(\boldsymbol{\theta})= (\mathbf{x} - \mathbf{s}(\boldsymbol{\theta}))^T (\mathbf{x} - \mathbf{s}(\boldsymbol{\theta})) = \sum_{n=0}^{N-1} (\mathbf{x}[n] - s[n;\boldsymbol{\theta}])^2$$

c. Optimality/Error Criterion

None In General

d. Performance

Depends on PDF of \mathbf{w} -no general formula is available

e. Comments

The fact that we are **minimizing a LS error** **minimizing a LS error criterion does not in general translate into minimizing the estimation error**. Also \mathbf{w} is a Gaussian RV $N(0,\sigma^2\mathbf{I})$, then the LSE is equivalent to MLE.

Method of Moments

a. Data Model/Assumptions

There are p moments $\mu_i = E(x^i[n])$ for $i = 1, 2, \dots, p$, which depends upon θ in a known way.
The entire PDF need not be known.

b. Estimator

If $\mu = \mathbf{h}(\theta)$, where \mathbf{h} is an invertible p -D function of θ & $\mu = [\mu_1 \mu_2 \dots \mu_p]^T$, then

$$\hat{\theta} = \mathbf{h}^{-1}(\hat{\mu})$$

c. Optimality/Error Criterion

None in General

d. Performance

For finite N it depends on PDF of \mathbf{x} . However for large data records (asymptotically), if

$\hat{\theta}_i = g_i(\hat{\mu})$ then,

$$E(\hat{\theta}_i) = g_i(\mu)$$

$$\text{var}(\hat{\theta}_i) = \left(\frac{\partial g_i}{\partial \hat{\mu}} \right)_{\hat{\mu}=\mu}^T \mathbf{C}_{\hat{\mu}} \left(\frac{\partial g_i}{\partial \hat{\mu}} \right)_{\hat{\mu}=\mu} \quad \text{for } i = 1, 2, \dots, p$$

e. Comments

Usually very ***easy to implement.***

Minimum Mean Square Error Estimator

a. Data model/Assumptions

The joint PDF $p(\mathbf{x}, \boldsymbol{\theta})$ is known, where $\boldsymbol{\theta}$ is considered to be a RV. Usually, $p(\mathbf{x}|\boldsymbol{\theta})$ is specified as the data model and $p(\boldsymbol{\theta})$ as the prior PDF for $\boldsymbol{\theta}$, so that $p(\mathbf{x}, \boldsymbol{\theta}) = p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})$.

b. Estimator

$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}|\mathbf{x})$, where the expectation represents the posterior PDF

$$p(\boldsymbol{\theta}|\mathbf{x}) = \frac{p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{\int p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})d\boldsymbol{\theta}}$$

If $\mathbf{x}, \boldsymbol{\theta}$ are jointly Gaussian, $\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}(\mathbf{x} - E(\mathbf{x}))$

c. Optimality/Error criterion

$\hat{\boldsymbol{\theta}}_i$ minimizes the Bayesian MSE

$$Bmse(\hat{\boldsymbol{\theta}}_i) = E[(\theta_i - \hat{\boldsymbol{\theta}}_i)^2] \quad i=1, 2, \dots, p \quad \text{where the Expectation is w.r.t } p(\mathbf{x}, \theta_i)$$

d. Performance

The error $\varepsilon_i = \theta_i - \hat{\boldsymbol{\theta}}_i$ has zero mean & variance

$$\text{var}(\varepsilon_i) = Bmse(\hat{\boldsymbol{\theta}}_i) = \int [\mathbf{C}_{\theta_i\mathbf{x}}]_{ii} p(\mathbf{x}) d\mathbf{x}$$

If $\mathbf{x}, \boldsymbol{\theta}$ are jointly Gaussian, then the error is Gaussian with zero mean & variance

$$\text{var}(\varepsilon_i) = [\mathbf{C}_{\boldsymbol{\theta}\boldsymbol{\theta}} - \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}\mathbf{C}_{\mathbf{x}\boldsymbol{\theta}}]_{ii}$$

e. Comments

In the non-Gaussian case, this will be difficult to implement.

Maximum A Posteriori Estimator

a. Data Model/Assumptions

Same as for the MMSE estimator.

b. Estimator

$\hat{\boldsymbol{\theta}}$ is the value of $\boldsymbol{\theta}$ that maximizes $p(\boldsymbol{\theta}|\mathbf{x})$ or, equivalently the value that maximizes $p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})$. If $\mathbf{x},\boldsymbol{\theta}$ are jointly Gaussian, $\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}(\mathbf{x} - E(\mathbf{x}))$

c. Optimality/Error Criterion

Minimizes the "**hit or miss**" cost function.

d. Performance

Depends on PDF-no general formula available. If $\mathbf{x},\boldsymbol{\theta}$ are jointly Gaussian, then the performance is identical to that of MMSE estimator.

e. Comments

For PDF's whose **mean & mode are the same**, the MMSE & MAP estimators will be identical, i.e., the Gaussian PDF, for example.

Linear Minimum Mean Square Error Estimator

a. Data Model/Assumptions

The first two moments of the joint PDF $p(\mathbf{x}, \boldsymbol{\theta})$ are known or the mean and covariance

$$\begin{bmatrix} E(\boldsymbol{\theta}) \\ E(\mathbf{x}) \end{bmatrix} \quad \begin{pmatrix} C_{\boldsymbol{\theta}\boldsymbol{\theta}} & C_{\boldsymbol{\theta}\mathbf{x}} \\ C_{\mathbf{x}\boldsymbol{\theta}} & C_{\mathbf{x}\mathbf{x}} \end{pmatrix}$$

b. Estimator

$$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}} \mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

c. Optimality/Error Criterion

$\hat{\boldsymbol{\theta}}_i$ has the minimum Bayesian MSE of all estimators that are *linear* function of \mathbf{x}

d. Performance

The error $\boldsymbol{\varepsilon}_i = \boldsymbol{\theta}_i - \hat{\boldsymbol{\theta}}_i$ has zero mean & variance

$$\text{var}(\boldsymbol{\varepsilon}_i) = \text{Bmse}(\hat{\boldsymbol{\theta}}_i) = [\mathbf{C}_{\boldsymbol{\theta}\boldsymbol{\theta}} - \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}} \mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1} \mathbf{C}_{\mathbf{x}\boldsymbol{\theta}}]_{ii}$$

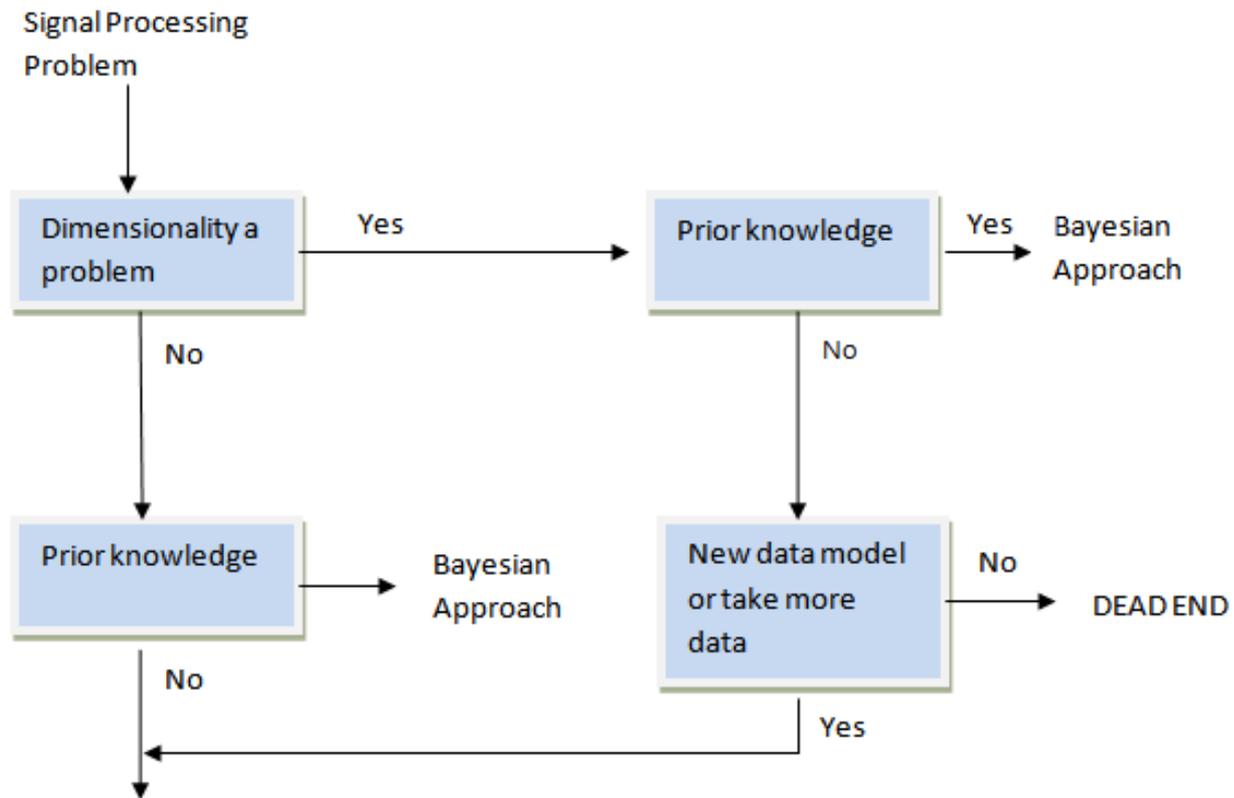
e. Comments

If $\mathbf{x}, \boldsymbol{\theta}$ are jointly Gaussian, this is identical to the MMSE & MAP estimators.

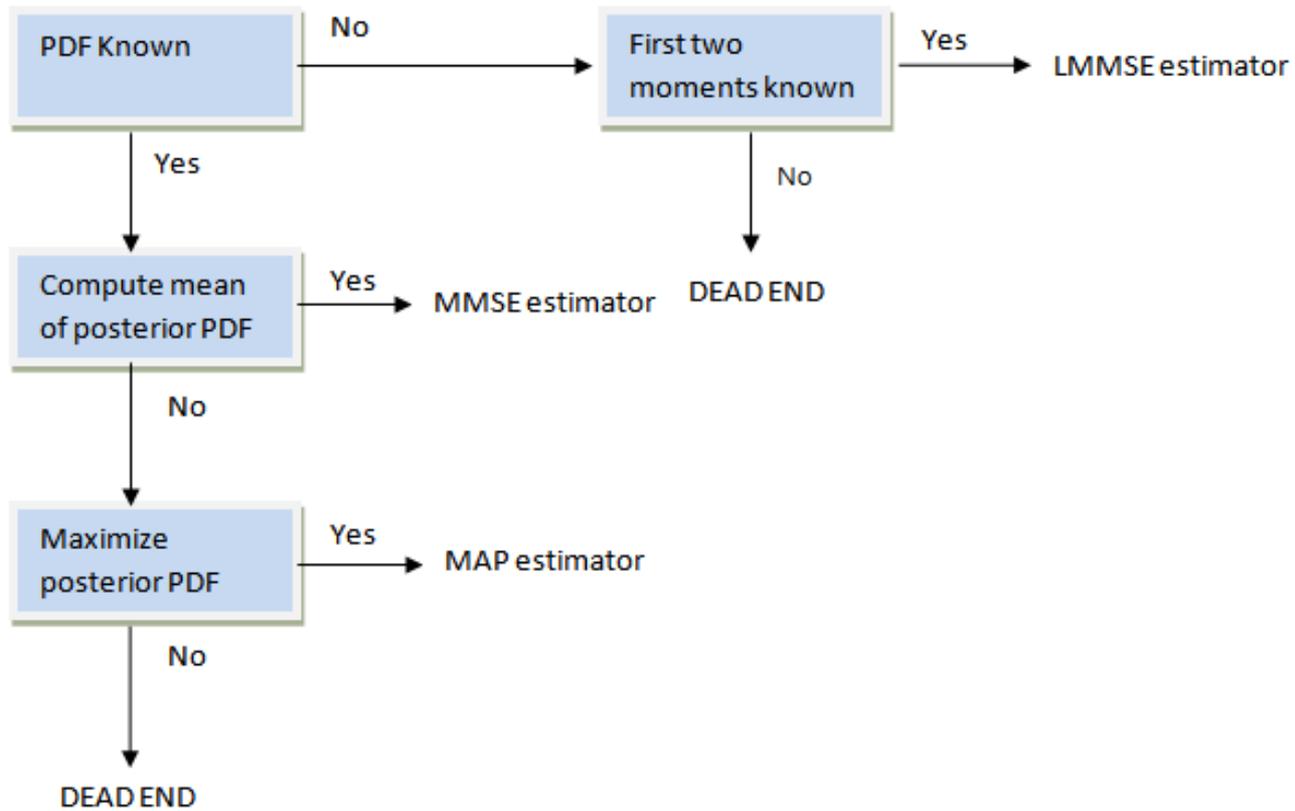
Choosing an Estimator

- The Choice of an estimator should begin with the search for an optimal estimator that is computationally feasible.
- If the search proves futile, then suboptimal estimators should be investigated.
- Even if the estimator is optimal for the given data model, its performance may not be adequate. In such a case, the data model may need to be modified.

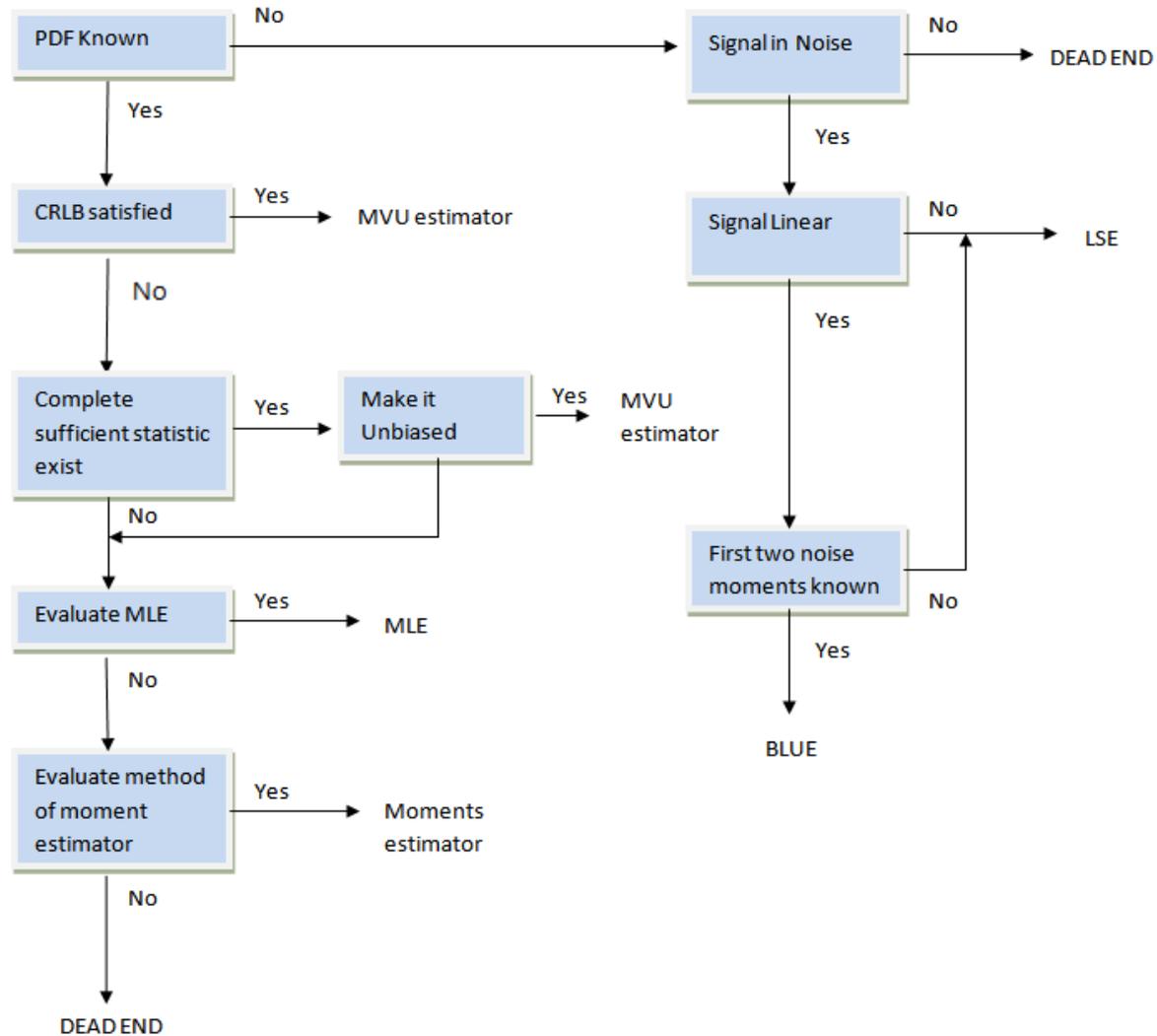
Classic v/s Bayesian



Bayesian Approach



Classical Approach



Final Comments

- Selection of Good Data Model
 - Complex enough to describe principle features
 - Simple enough to allow an estimator that is optimal & easily feasible.
- Determination of optimal estimators not possible
 - Search of an MVU estimator in classical approach
- Implementation of optimal estimator not possible
 - MMSE estimator in Bayesian estimation