

# **EE 602 Term Paper on**

# **Summary of Estimators**



**Submitted by:**

**Vishesh lokras(Y8104067)**

**Avadhesh Kr Yadav(Y8104010)**

# Introduction

**There are two methods to estimate a parameter vector  $\theta$**

First method is Classical Approach in which the Unknown parameter vector  $\theta$  is a Deterministic Constant. Various Estimators in classical Approach are given below:

- Cramer-Rao Lower Bound
- Rao-Blackwell-Lehmann-Scheffe
- Best Linear Unbiased Estimator
- Maximum Likelihood Estimator
- Least Squares Estimator
- Method of Moments

Second method is Bayesian Approach in which the Known parameter vector  $\theta$  is Realization of a Random Vector. Various Estimators in Bayesian Approach are given below:

- Minimum Mean Square Error Estimator
- Maximum A Posteriori Estimator
- Linear Minimum Mean Square Error Estimator

# Cramer Rao Lower Bound(CRLB)

## a. Data Model/Assumption

PDF  $p(\mathbf{x};\boldsymbol{\theta})$  is known.

## b. Estimator

If the equality condition for the CRLB

$$\frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{I}(\boldsymbol{\theta})(\mathbf{g}(\mathbf{x}) - \boldsymbol{\theta})$$

is satisfied, then the estimator is

$$\hat{\boldsymbol{\theta}} = \mathbf{g}(\mathbf{x})$$

where  $\mathbf{I}(\boldsymbol{\theta})$  is a  $p \times p$  matrix dependent only on  $\boldsymbol{\theta}$  and  $\mathbf{g}(\mathbf{x})$  is a  $p$ -D function of data  $\mathbf{x}$ .

## c. Optimality/Criterion

$\hat{\boldsymbol{\theta}}$  achieves the CRLB, the lower bound on the variance for any unbiased estimator (and hence is said to be efficient), and is therefore the MVU estimator.

## d. Performance

It is **unbiased** and has the **minimum variance**

$$\text{var}(\hat{\theta}_i) = [\mathbf{I}^{-1}(\boldsymbol{\theta})]_{ii} \quad i=1,2,\dots,p$$

Where

$$[\mathbf{I}(\boldsymbol{\theta})]_{ij} = \mathbb{E} \left[ \frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \theta_i} \frac{\partial \ln p(\mathbf{x}; \boldsymbol{\theta})}{\partial \theta_j} \right]$$

## e. Comments

An efficient **Estimator may not exist** & this approach may fail

# Rao-Blackwell-Lehmann-Scheffe

## *a. Data Model/Assumption*

PDF  $p(\mathbf{x};\boldsymbol{\theta})$  is known.

## *b. Estimator*

- i. Find a *sufficient statistic*  $\mathbf{T}(\mathbf{x})$  by factoring PDF as

$$p(\mathbf{x};\boldsymbol{\theta})=g(\mathbf{T}(\mathbf{x}),\boldsymbol{\theta})h(\mathbf{x})$$

- ii. If  $E[\mathbf{T}(\mathbf{x})]=\boldsymbol{\theta}$ , then  $\hat{\boldsymbol{\theta}} = \mathbf{T}(\mathbf{x})$ . If not, we must find a  $p$ -dimensional function  $\mathbf{g}$  so that  $E[\mathbf{g}(\mathbf{T})]=\boldsymbol{\theta}$ , and then  $\hat{\boldsymbol{\theta}} = \mathbf{g}(\mathbf{T})$ .

## *c. Optimality/Criterion*

$\hat{\boldsymbol{\theta}}$  is the MVU estimator.

## *d. Performance*

**Unbiased** but **variance depends upon PDF** -no general formula is available

## *e. Comments*

“**Completeness**” of sufficient statistic must be checked. A  $p$ -dimensional sufficient statistic may not exist so this method may fail.

# Best Linear Unbiased Estimator

## a. Data Model/Assumptions

$$E[\mathbf{x}] = \mathbf{H}\boldsymbol{\theta}$$

where  $\mathbf{H}$  is an  $N \times p$  ( $N > p$ ) known matrix and  $\mathbf{C}$ , the covariance matrix of  $\mathbf{x}$ , is known. Equivalently, we have

$$\mathbf{x} = \mathbf{H}\boldsymbol{\theta} + \mathbf{w}$$

where  $E(\mathbf{w}) = 0$  and  $\mathbf{C}_w = \mathbf{C}$ .

## b. Estimator

$$\hat{\boldsymbol{\theta}} = (\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1} \mathbf{H}^T \mathbf{C}^{-1} \mathbf{x}$$

## c. Optimality/Error Criterion

$\hat{\boldsymbol{\theta}}_i$  for  $i = 1, 2, \dots, p$  has **minimum variance** of all unbiased estimators that are linear in  $\mathbf{x}$ .

## d. Performance

$\hat{\boldsymbol{\theta}}_i$  for  $i = 1, 2, \dots, p$  is unbiased. The variance is

$$\text{var}(\boldsymbol{\theta}_i) = [(\mathbf{H}^T \mathbf{C}^{-1} \mathbf{H})^{-1}]_{ii} \quad i = 1, 2, \dots, p$$

## e. Comments

If  $\mathbf{w}$  is a Gaussian RV, then is  $\hat{\boldsymbol{\theta}}$  also the **MVU estimator** (for all linear & non linear functions of  $\mathbf{x}$ ).

# Maximum Likelihood Estimator

a. Data Model/Assumptions

PDF  $p(\mathbf{x};\boldsymbol{\theta})$  is known

b. Estimator

$\hat{\boldsymbol{\theta}}$  is the value of  $\boldsymbol{\theta}$  maximizing  $p(\mathbf{x};\boldsymbol{\theta})$ , where  $\mathbf{x}$  is replaced by the observed data samples.

c. Optimality/Error Criterion

Not optimal in general. Under certain conditions on the PDF, however, the MLE is efficient for large data records or as  $N \rightarrow \infty$  (asymptotically). Hence, asymptotically it is the MVU estimator.

d. Performance

For finite  $N$  depends on PDF-no general formula is available. Asymptotically, under certain conditions

$$\hat{\boldsymbol{\theta}} = N(\boldsymbol{\theta}, I^{-1}(\boldsymbol{\theta}))$$

e. Comments

If an MVU estimator exists, the MLE procedure will produce it .

# Least Squares Estimator

## *a. Data Model/Assumption*

$$x[n]=s[n;\boldsymbol{\theta}] + w[n] \quad n=0,1,\dots,N-1$$

where the signal  $s[n;\boldsymbol{\theta}]$  depends explicitly on the unknown parameters. Equivalently, the model is

$$\mathbf{x}=\mathbf{s}(\boldsymbol{\theta}) + \mathbf{w}$$

where  $\mathbf{s}$  is a known  $N$ -dimensional function of  $\boldsymbol{\theta}$  and the noise or perturbation  $\mathbf{w}$  has zero mean.

## *b. Estimator*

$\hat{\boldsymbol{\theta}}$  is the value that minimizes

$$J(\boldsymbol{\theta}) = (\mathbf{x} - \mathbf{s}(\boldsymbol{\theta}))^T(\mathbf{x} - \mathbf{s}(\boldsymbol{\theta})) = \sum_{n=0}^{N-1} (x[n] - s[n; \boldsymbol{\theta}])^2$$

## *c. Optimality/Error Criterion*

None In General

## *d. Performance*

Depends on PDF of  $\mathbf{w}$ -no general formula is available

## *e. Comments*

The fact that we are *minimizing a LS error minimizing a LS error criterion does not in general translate into minimizing the estimation error*. Also  $\mathbf{w}$  is a Gaussian RV  $N(0,\sigma^2\mathbf{I})$ , then the LSE is equivalent to MLE.

# Method of Moments

## a. Data Model/Assumptions

There are  $p$  moments  $\mu_i = E(x^i[n])$  for  $i = 1, 2, \dots, p$ , which depends upon  $\theta$  in a known way. *The entire PDF need not be known.*

## b. Estimator

If  $\mu = h(\theta)$ , where  $h$  is an invertible  $p$ -D function of  $\theta$  &  $\mu = [\mu_1 \mu_2 \dots \mu_p]^T$ , then

$$\hat{\theta} = h^{-1}(\hat{\mu})$$

## c. Optimality/Error Criterion

None in General

## d. Performance

For finite  $N$  it depends on PDF of  $\mathbf{x}$ . However for large data records (asymptotically), if

$$E(\hat{\theta}_i) = g_i(\mu)$$

$$\text{var}(\hat{\theta}_i) = \left( \frac{\partial g_i}{\partial \hat{\mu}} \right)_{\hat{\mu}=\mu}^T C_{\hat{\mu}} \left( \frac{\partial g_i}{\partial \hat{\mu}} \right)_{\hat{\mu}=\mu}$$

for  $i = 1, 2, \dots, p$

## e. Comments

Usually very ***easy to implement.***



# Minimum Mean Square Error Estimator

## a. Data model/Assumptions

The joint PDF  $p(\mathbf{x}, \boldsymbol{\theta})$  is known, where  $\boldsymbol{\theta}$  is considered to be a RV. Usually,  $p(\mathbf{x}|\boldsymbol{\theta})$  is specified as the data model and  $p(\boldsymbol{\theta})$  as the prior PDF for  $\boldsymbol{\theta}$ , so that  $p(\mathbf{x}, \boldsymbol{\theta}) = p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})$ .

## b. Estimator

$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}|\mathbf{x})$ , where the expectation represents the posterior PDF

$$p(\boldsymbol{\theta}|\mathbf{x}) = \frac{p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{\int p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})d\boldsymbol{\theta}}$$

If  $\mathbf{x}, \boldsymbol{\theta}$  are jointly Gaussian,  $\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}(\mathbf{x} - E(\mathbf{x}))$

## c. Optimality/Error criterion

$\hat{\boldsymbol{\theta}}_i$  minimizes the Bayesian MSE

$Bmse(\hat{\boldsymbol{\theta}}_i) = E[(\theta_i - \hat{\boldsymbol{\theta}}_i)^2]$   $i=1, 2, \dots, p$  where the Expectation is w.r.t  $p(\mathbf{x}, \theta_i)$

## d. Performance

The error  $\varepsilon_i = \theta_i - \hat{\boldsymbol{\theta}}_i$  has zero mean & variance

$$\text{var}(\varepsilon_i) = Bmse(\hat{\boldsymbol{\theta}}_i) = \int [\mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}]_{ii} p(\mathbf{x}) d\mathbf{x}$$

If  $\mathbf{x}, \boldsymbol{\theta}$  are jointly Gaussian, then the error is Gaussian with zero mean & variance

$$\text{var}(\varepsilon_i) = [\mathbf{C}_{\boldsymbol{\theta}\boldsymbol{\theta}} - \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}\mathbf{C}_{\mathbf{x}\boldsymbol{\theta}}]_{ii}$$

# Maximum A Posteriori Estimator

## a. *Data Model/Assumptions*

Same as for the MMSE estimator.

## b. *Estimator*

$\hat{\boldsymbol{\theta}}$  is the value of  $\boldsymbol{\theta}$  that maximizes  $p(\boldsymbol{\theta}|\mathbf{x})$  or, equivalently the value that maximizes  $p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})$ . If  $\mathbf{x},\boldsymbol{\theta}$  are jointly Gaussian,

$$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + \mathbf{C}_{\boldsymbol{\theta}\mathbf{x}}\mathbf{C}_{\mathbf{x}\mathbf{x}}^{-1}(\mathbf{x} - E(\mathbf{x}))$$

## c. *Optimality/Error Criterion*

Minimizes the "**hit or miss**" cost function.

## d. *Performance*

Depends on PDF-no general formula available. If  $\mathbf{x},\boldsymbol{\theta}$  are jointly Gaussian, then the performance is identical to that of MMSE estimator.

## e. *Comments*

For PDF's whose **mean & mode are the same**, the MMSE & MAP estimators will be identical, i.e., the Gaussian PDF, for example.

# Linear Minimum Mean Square Error Estimator

## a. Data Model/Assumptions

The first two moments of the joint PDF  $p(\mathbf{x}, \boldsymbol{\theta})$  are known or the mean and covariance

$$\begin{bmatrix} E(\boldsymbol{\theta}) \\ E(\mathbf{x}) \end{bmatrix} = \begin{pmatrix} C_{\boldsymbol{\theta}\boldsymbol{\theta}} & C_{\boldsymbol{\theta}\mathbf{x}} \\ C_{\mathbf{x}\boldsymbol{\theta}} & C_{\mathbf{x}\mathbf{x}} \end{pmatrix}$$

## b. Estimator

$$\hat{\boldsymbol{\theta}} = E(\boldsymbol{\theta}) + C_{\boldsymbol{\theta}\mathbf{x}} C_{\mathbf{x}\mathbf{x}}^{-1} (\mathbf{x} - E(\mathbf{x}))$$

## c. Optimality/Error Criterion

$\hat{\boldsymbol{\theta}}_i$  has the minimum Bayesian MSE of all estimators that are *linear* function of  $\mathbf{x}$

## d. Performance

The error  $\varepsilon_i = \theta_i - \hat{\boldsymbol{\theta}}_i$  has zero mean & variance

$$\text{var}(\varepsilon_i) = \text{Bmse}(\hat{\boldsymbol{\theta}}_i) = [C_{\boldsymbol{\theta}\boldsymbol{\theta}} - C_{\boldsymbol{\theta}\mathbf{x}} C_{\mathbf{x}\mathbf{x}}^{-1} C_{\mathbf{x}\boldsymbol{\theta}}]_{ii}$$

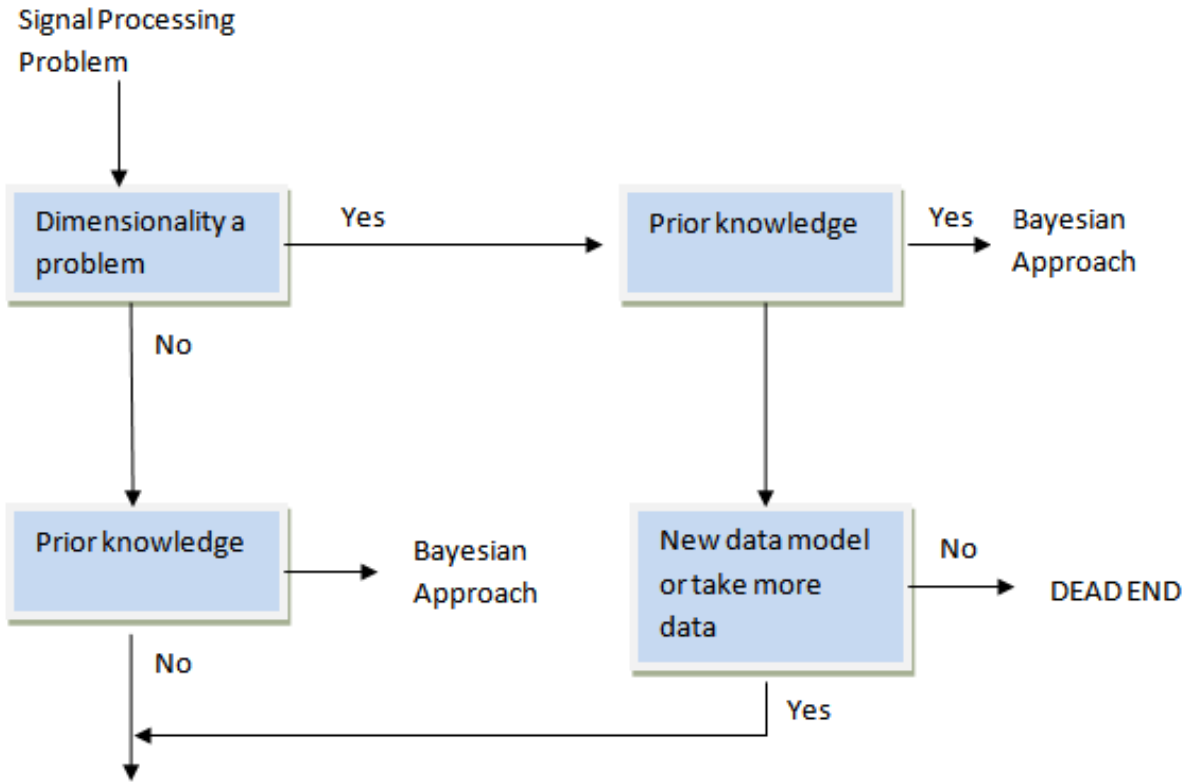
## e. Comments

If  $\mathbf{x}, \boldsymbol{\theta}$  are jointly Gaussian, this is identical to the MMSE & MAP estimators.

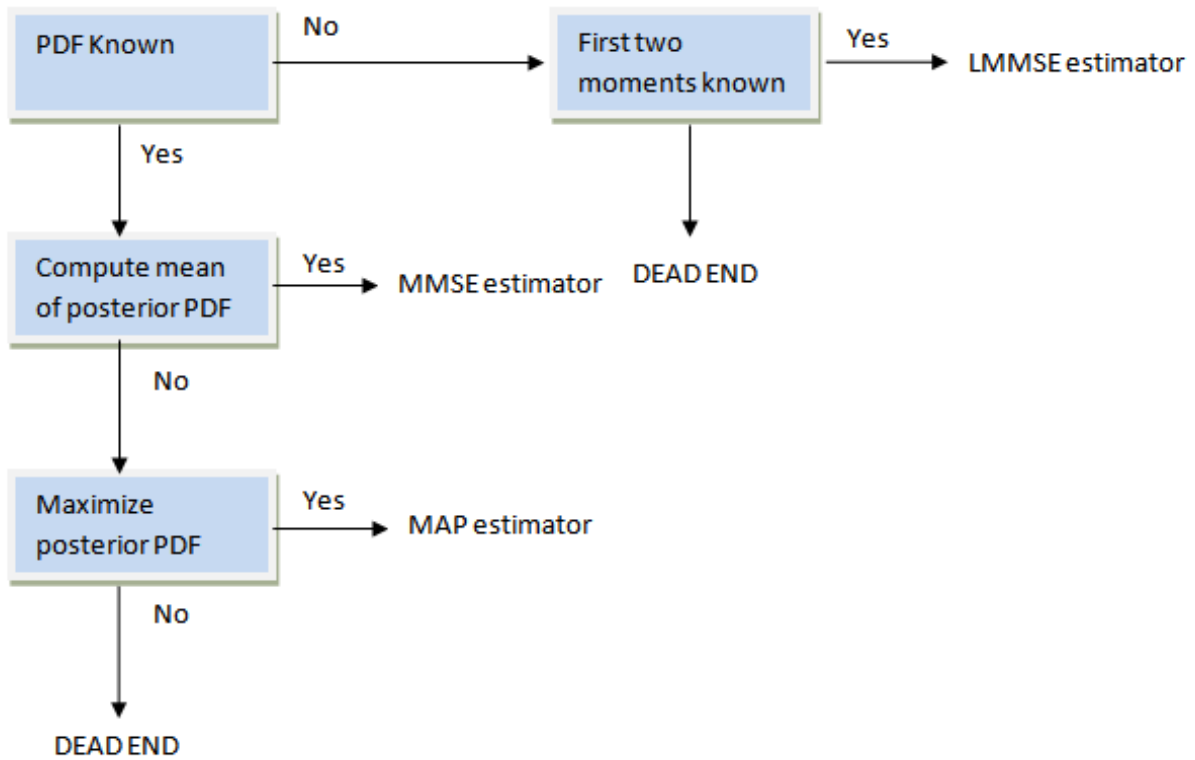
# Choosing an Estimator

- The Choice of an estimator should begin with the search for an optimal estimator that is computationally feasible.
- If the search proves futile, then suboptimal estimators should be investigated.
- Even if the estimator is optimal for the given data model, its performance may not be adequate. In such a case, the data model may need to be modified.

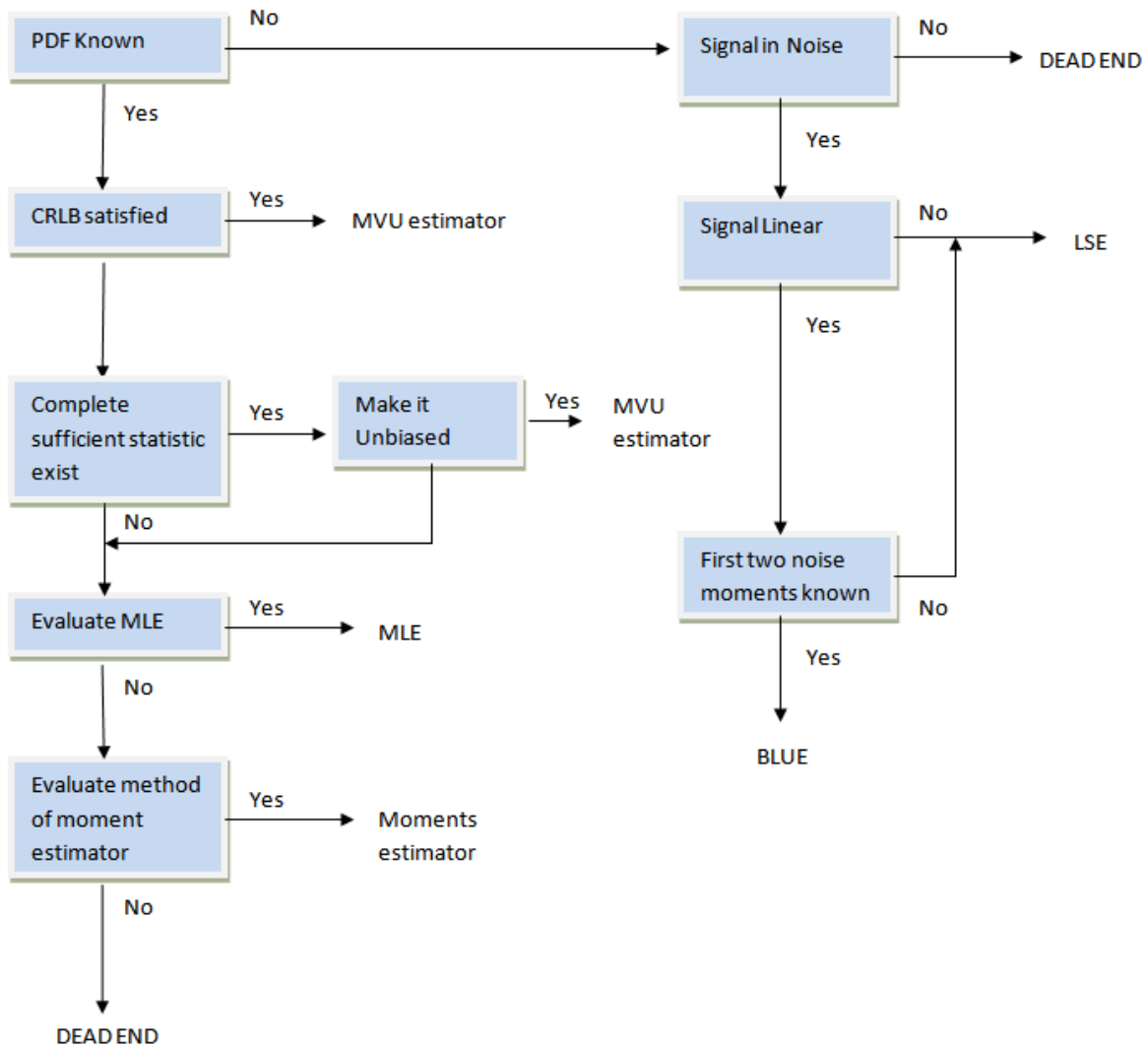
# Classic v/s Bayesian



# Bayesian Approach



# Classical Approach



# Final Comments

- Selection of Good Data Model
  - Complex enough to describe principle features
  - Simple enough to allow an estimator that is optimal & easily feasible.
- Determination of optimal estimators not possible
  - Search of an MVU estimator in classical approach
- Implementation of optimal estimator not possible
  - MMSE estimator in Bayesian estimation.