EE-602

STATISTICAL SIGNAL PROCESSING

TERM PAPER REPORT ON

CONSTRAINED AND NON-LINEAR LEAST SQUARES



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CONSTRAINED LEAST SQUARES

At times we come across the LS problems where unknown parameters must be constrained. Assuming parameter θ is subject to r<p independent constraints, then we summarize the constraints as

$$A\theta = b$$

Where A is a r x p matrix and b is a known r x 1 vector.

For instance, let p=2 and one parameter is negative of other, then the constraint would be

$$\theta_1 + \theta_2 = 0$$

Then we have

 $A=\begin{bmatrix} 1 & 1 \end{bmatrix}$ and b=0

The condition for the constraints to be independent is that the matrix A should be full rank (equal to r). There are only p-r independent parameters.

To find the Least Squares Estimator (LSE) subject to linear constraints, we use Lagrangian multiplier.

We determine θ_{C} (c denotes the constrained LSE) by minimizing the Lagrangian

$$J_{C} = (x-H\theta)^{T}(x-H\theta) + \lambda^{T} (A\theta-b)$$

Where λ is a r x 1 vector of Lagrangian multipliers.

Expanding above expression, we have

 $J_{C} = x^{T}x - 2\theta^{T}H^{T}x + \theta^{T}H^{T}H\theta + \lambda^{T}A\theta - \lambda^{T}b$

Taking the gradient w.r.t. θ , we have

$$\frac{\mathrm{d}J}{\mathrm{d}\theta} c = -2H^{\mathrm{T}}x + 2H^{\mathrm{T}}H\theta + A^{\mathrm{T}}\lambda$$

Setting it to zero, we have

$$\hat{\theta}_{C} = (H^{T}H)^{-1}H^{T}x^{-}\frac{1}{2}(H^{T}H)^{-1}A^{T}\lambda$$
$$= \hat{\theta} - \frac{1}{2}(H^{T}H) - 1A^{T}\lambda \qquad (i)$$

Where $\,\widehat{\theta}\,$ is the unconstrained LSE.

Now, to find $\boldsymbol{\lambda},$ we impose the constrain

 $A\theta = b$

So that

$$A\theta_{C} = A\widehat{\theta} - \frac{1}{2}A(H^{T}H)^{-}1A^{T}\lambda = b$$

Hence

$$\frac{1}{2}\lambda = [A(H^{T}H)^{-1}A^{T}]^{-1}(A\theta-b)$$

Substituting in equation (i) gives

$$\hat{\theta}_{C} = \hat{\theta} \cdot (H^{T}H)^{-1}A^{T}[A(H^{T}H)^{-1}A^{T}]^{-1}(A\hat{\theta} \cdot b)$$
(ii)

Where

$$\hat{\theta} = (\mathbf{H}^{\mathrm{T}}\mathbf{H})^{-1}\mathbf{H}^{\mathrm{T}}\mathbf{x}$$

EXAMPLE OF CONSTRAINED LEAST SQUARES

If the signal model is

$$\mathbf{s}[\mathbf{n}] = \begin{cases} \theta 1\\ \theta 2\\ 0 \end{cases}$$

and we observe {x[0],x[1],x[2]}, then the observation matrix is

$$\mathbf{H} = \begin{bmatrix} 1 & 0\\ 0 & 1\\ 0 & 0 \end{bmatrix}$$

The signal vector is

$$\mathbf{s} = \mathbf{H}\boldsymbol{\theta} = \begin{bmatrix} \boldsymbol{\theta} 1 \\ \boldsymbol{\theta} 2 \\ \mathbf{0} \end{bmatrix}$$

The unconstrained LSE is

$$\hat{\theta} = (\mathbf{H}^{\mathrm{T}}\mathbf{H})^{-1}\mathbf{H}^{\mathrm{T}}\mathbf{x} = \begin{bmatrix} \mathbf{x}[\mathbf{0}]\\ \mathbf{x}[\mathbf{1}] \end{bmatrix}$$

and the signal estimate is

$$\hat{\mathbf{s}} = \mathbf{H}\hat{\boldsymbol{\theta}} = \begin{bmatrix} \mathbf{x}[0] \\ \mathbf{x}[1] \\ \mathbf{0} \end{bmatrix}$$

Now assume that we know a priori that

$$\theta_1 = \theta_2$$

thus we have

$$[1 -1] \theta = 0$$

so that

$$A = \begin{bmatrix} 1 & -1 \end{bmatrix}$$
 and $b = 0$

Now noting that

 $\mathbf{H}^{\mathrm{T}}\mathbf{H} = \mathbf{I}$

so from equation (ii) we have

$$\hat{\theta}_{C} = \hat{\theta} - A^{T} [AA^{T}]^{-1} A$$
$$= [I - A^{T} [AA^{T}]^{-1} A] \hat{\theta}$$

on solving the above expression, we have

$$\hat{\theta}_{C} = \begin{bmatrix} \frac{1}{2} (x[0] + x[1]) \\ \frac{1}{2} (x[0] + x[1]) \end{bmatrix}$$

and the constrained signal becomes

$$\hat{s}_{c} = H\hat{\theta}_{c} = \begin{bmatrix} \frac{1}{2}(x[0] + x[1]) \\ \frac{1}{2}(x[0] + x[1]) \\ 0 \end{bmatrix}$$

NONLINEAR LEAST SQUARES

Least Square Procedure that estimates model parameters $\boldsymbol{\theta}$ by minimizing the Least Square Error Criterion:

$$J = (x-s(\theta))^{T} (x-s(\theta))$$

 $s(\theta) = Signal \mod for \mathbf{x}$, with its dependence on θ .

In the Linear Least Square problem the signal takes on special form $s(\theta) = H\theta$, which leads to the simple linear Least Square Estimator. In general, $s(\theta)$ can't be expressed in this manner but is an N-dimensional nonlinear function of θ . In such a case the minimization of J becomes much more difficult, if not impossible. This type of nonlinear Least Square problem is often termed a *nonlinear regression problem*.

Methods that can reduce complexity of Problem for determining nonlinear LSEs:

- 1. Transformation of parameters
- 2. Separability of parameters

1. <u>Transformation of parameters</u>

One-to-one transformation of $\boldsymbol{\theta}$ that produces a linear signal model in the new space

$$\alpha = g(\theta)$$

g: p-dimensional function of θ whose inverse exists. If a **g** can be found so that

 $s(\theta) = s(g^{-1}(\alpha)) = H\alpha$

then the signal model will be linear in α .

Thus the **nonlinear** LSE of $\boldsymbol{\theta}$ by

$$\widehat{\theta} = g^{-1}(\widehat{\alpha})$$

where: $\widehat{\alpha} = (H^T H)^{-1} H^T x$

Approach: minimization can be carried out in any transformed space that is obtained by a one-to-one mapping & then converted back to original space.

2. <u>Separability of parameters</u>

A separable signal model has the form

$$s = H(\alpha)\beta$$
$$\theta = \begin{bmatrix} \alpha \\ \beta \end{bmatrix} = \begin{bmatrix} (p-q) \times 1 \\ q \times 1 \end{bmatrix}$$

 $H(\alpha)$ is an $N \times q$ matrix dependent on α .

This model is linear in β but nonlinear in α . The LS Error may be minimized with respect to β and thus reduced to a function of α only.

Since

$$J(\alpha,\beta) = (x - H(\alpha)\beta)^{T} (x - H(\alpha)\beta)$$

the $\boldsymbol{\beta}$ that minimizes J for a given a is

 $\hat{\boldsymbol{\beta}} = (\boldsymbol{H}^{\mathrm{T}}(\alpha)\boldsymbol{H}(\alpha))^{-1} \boldsymbol{H}^{\mathrm{T}}(\alpha)\boldsymbol{x}$

And the resulting LS error is

 $J(\alpha, \hat{\beta}) = x^{T} [I - H(\alpha) (H^{T}(\alpha) H(\alpha))^{-1} H(\alpha)^{T}] x$

The Problem now reduces to a maximization of

$$x^{T} H(\alpha)(H^{T}(\alpha)H(\alpha))^{-1} H(\alpha)^{T} x$$

over α . If, for instance, q=p-1, so that α is a scalar, then a grid can possibly be used. This should be contrasted with the original minimization of a *p*-dimensional function.

Example: Sinusoidal Parameter Estimation

For a sinusoidal signal model

$$s[n] = A\cos(2\pi f_o n + \phi)$$

It is desired to estimate the amplitude A, where A > 0, phase ϕ .

The frequency f_o is assumed known.

LSE is obtained by minimizing

$$J = \sum_{n=0}^{N-1} (x[n] - A\cos(2\pi f_o n + \phi))^2$$

Over A & $\boldsymbol{\varphi},$ a nonlinear LS problem. However, because

 $A\cos(2\pi f_0 n + \phi) = A\cos(\phi)\cos(2\pi f_0 n) - A\sin(\phi)\sin(2\pi f_0 n)$

If we let $\alpha_1 = A \cos(\phi)$

 $\alpha_2 = -A \sin(\phi)$

Then the signal model becomes

$$s[n] = \alpha_1 \cos(2\pi f_0 n) + \alpha_2 \sin(2\pi f_0 n)$$

In the Matrix form this is $s = H\alpha$

Where,

$$H = \begin{bmatrix} 1 & 0\\ \cos(2\pi f_0 n) & \sin(2\pi f_0 n)\\ \vdots & \vdots\\ \cos(2\pi f_0 (N-1)) & \sin(2\pi f_0 (N-1)) \end{bmatrix}$$

which is now linear in the terms of new parameters .

The LSE of $\alpha\,$ is

$$\hat{\alpha} = (\mathbf{H}^{\mathrm{T}}\mathbf{H})^{-1}\mathbf{H}^{\mathrm{T}}\mathbf{x}$$

and to find $\hat{\theta}$ we must find the inverse transformation g⁻¹(α).

This is

$$A = (\alpha_1^2 + \alpha_2^2)^{1/2}$$
$$\phi = \tan^{-1} \left(\frac{\mathbb{P} - \alpha_2^2}{\alpha_1} \right)$$

So that the nonlinear LSE for this problem is given by

$$\widehat{\theta} = \begin{bmatrix} \widehat{A} \\ \widehat{\phi} \end{bmatrix}$$
$$\widehat{\theta} = \begin{bmatrix} \sqrt{\alpha_1^2 + \alpha_2^2} \\ \tan^{-1} \begin{bmatrix} \frac{2}{\alpha_1} \\ \frac{2}{\alpha_1} \end{bmatrix}$$

Where: $\widehat{\alpha} = (H^T H)^{-1} H^T x$